## Supplementary appendix

## SA1.1 Additional figures and tables

Table SA2: Tests of differences between labor market histories above and below the median wage

|  |  | Days in part-time over the last 5 years |  | Days in nonemployment over the last 5 years |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Year | 1985 | 2010 | 1985 | 2010 |
| Males | P -value of difference in means | 0.0035 | 0 | 0 | 0 |
|  | P -value of difference in variances | 0.0067 | 0 | 0 | 0 |
| Females | P -value of difference in means | 0 | 0 | 0 | 0 |
|  | P -value of difference in variances | 0 | 0 | 0 | 0 |

Figure SA2: Share of occupation categories


Table SA3: Descriptives of combined full-time and part-time samples

| Males |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | 1985 |  |  | 2010 |  |
|  | mean | sd | mean | sd |  |
| Real wage in Euro | 69.89 | 47.51 | 81.61 | 48.15 |  |
| Log real wage | 4.15 | 0.40 | 4.27 | 0.52 |  |
| No/other degree indicator | 0.19 | 0.40 | 0.08 | 0.28 |  |
| Vocational degree indicator | 0.70 | 0.46 | 0.71 | 0.46 |  |
| University degree indicator | 0.07 | 0.25 | 0.16 | 0.36 |  |
| Work experience | 27.32 | 11.19 | 29.08 | 10.23 |  |
| No. of days in full time last 5 years | 1540.45 | 494.58 | 1494.81 | 544.03 |  |
| Fulltime spell in previous year? | 0.96 | 0.20 | 0.94 | 0.24 |  |
| No. of days in part time last 5 years | 6.29 | 78.41 | 37.03 | 203.05 |  |
| Part-time spell in previous year? | 0.01 | 0.08 | 0.03 | 0.18 |  |
| Agriculture and mining | 0.03 | 0.17 | 0.02 | 0.13 |  |
| Plastics, rubber, mineral products | 0.03 | 0.17 | 0.03 | 0.17 |  |
| Chemicals | 0.03 | 0.18 | 0.02 | 0.15 |  |
| Machinery and metal products | 0.15 | 0.36 | 0.13 | 0.33 |  |
| Transport- and electrical equipment | 0.12 | 0.32 | 0.10 | 0.31 |  |
| Food and basic consumption | 0.10 | 0.31 | 0.07 | 0.25 |  |
| Hotels and restaurants | 0.01 | 0.11 | 0.02 | 0.14 |  |
| Construction | 0.12 | 0.32 | 0.08 | 0.27 |  |
| Trade | 0.12 | 0.33 | 0.14 | 0.35 |  |
| Transport and communication | 0.06 | 0.24 | 0.07 | 0.26 |  |
| Financial and insurance | 0.08 | 0.27 | 0.18 | 0.38 |  |
| Public services | 0.04 | 0.20 | 0.05 | 0.21 |  |
| Health and Education | 0.04 | 0.18 | 0.06 | 0.23 |  |
| Public administration | 0.06 | 0.24 | 0.04 | 0.20 |  |


| Females | 1985 |  | 2010 |  |
| :--- | :---: | :---: | :---: | :---: |
|  | mean | sd | mean | sd |
| Real wage in Euro | 43.64 | 20.41 | 55.40 | 33.52 |
| Log real wage | 3.67 | 0.48 | 3.85 | 0.59 |
| No/other degree indicator | 0.28 | 0.45 | 0.08 | 0.28 |
| Vocational degree indicator | 0.65 | 0.48 | 0.73 | 0.44 |
| University degree indicator | 0.03 | 0.16 | 0.11 | 0.32 |
| Work experience | 24.99 | 11.90 | 28.62 | 10.99 |
| No. of days in full time last 5 years | 1199.98 | 696.74 | 1048.65 | 759.52 |
| Fulltime spell in previous year? | 0.81 | 0.39 | 0.72 | 0.45 |
| No. of days in part time last 5 years | 209.50 | 513.80 | 366.32 | 642.41 |
| Part-time spell in previous year? | 0.14 | 0.35 | 0.25 | 0.43 |
| Agriculture and mining | 0.01 | 0.08 | 0.01 | 0.08 |
| Plastics, rubber, mineral products | 0.02 | 0.14 | 0.01 | 0.10 |
| Chemicals | 0.02 | 0.14 | 0.01 | 0.12 |
| Machinery and metal products | 0.05 | 0.22 | 0.03 | 0.18 |
| Transport- and electrical equipment | 0.07 | 0.26 | 0.04 | 0.20 |
| Food and basic consumption | 0.12 | 0.33 | 0.06 | 0.23 |
| Hotels and restaurants | 0.03 | 0.18 | 0.03 | 0.18 |
| Construction | 0.02 | 0.14 | 0.02 | 0.13 |
| Trade | 0.19 | 0.39 | 0.16 | 0.37 |
| Transport and communication | 0.03 | 0.17 | 0.04 | 0.19 |
| Financial and insurance | 0.12 | 0.33 | 0.19 | 0.40 |
| Public services | 0.06 | 0.23 | 0.06 | 0.24 |
| Health and Education | 0.18 | 0.38 | 0.26 | 0.44 |
| Public administration | 0.08 | 0.28 | 0.07 | 0.26 |

Figure SA3: Inequality development base year 1985, specification EEHOI of total employment


Figure SA4: Inequality development base year 2010, specification EEHOI of total employment

Table SA4: Change in inequality measures since 1985, for males, composition adjusted to total employment in base year 2010

|  | Observed | Ed |  |  | Ed+Ex |  |  | Ed+Ex+Hist |  |  | Ed+Ex+Hist+Occ+Ind |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Increase | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment |
| 85/15 | 0.309 | 0.222 | 36.71\% | 36.71\% | 0.215 | 39.24\% | 2.53\% | 0.159 | 58.32\% | 19.08\% | 0.166 | 55.96\% | -2.36\% |
| 85/50 | 0.144 | 0.069 | 60.46\% | 60.46\% | 0.051 | 71.68\% | 11.22\% | 0.045 | 75.67\% | 3.99\% | 0.036 | 81.00\% | 5.33\% |
| 50/15 | 0.164 | 0.153 | 10.64\% | 10.64\% | 0.164 | 0.71\% | -9.93\% | 0.114 | 41.36\% | 40.65\% | 0.131 | 29.29\% | -12.07\% |
| 90/10 residual | 0.195 | 0.160 | -4.04\% | -4.04\% | 0.149 | 4.87\% | 8.91\% | 0.056 | 68.78\% | 63.91\% | 0.058 | 67.67\% | -1.11\% |

Table SA5: Change in inequality measures since 1985, for females, composition adjusted to total employment in base year 2010

|  | Observed | Ed |  |  | Ed+Ex |  |  | Ed+Ex+Hist |  |  | Ed+Ex+Hist+Occ+Ind |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Increase | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment |
| 85/15 | 0.234 | 0.173 | 31.10\% | 31.10\% | 0.087 | 68.30\% | 37.20\% | 0.035 | 88.04\% | 19.74\% | 0.062 | 78.15\% | -9.89\% |
| 85/50 | 0.098 | 0.047 | 58.46\% | 58.46\% | 0.047 | 58.46\% | 0.00\% | 0.031 | 73.55\% | 15.09\% | 0.039 | 66.00\% | -7.55\% |
| 50/15 | 0.136 | 0.127 | 8.88\% | 8.88\% | 0.041 | 75.07\% | 66.19\% | 0.004 | 97.73\% | 22.66\% | 0.023 | 86.39\% | -11.34\% |
| 90/10 | 0.158 | 0.156 | -15.97\% | -15.97\% | 0.077 | 47.55\% | 63.52\% | 0.036 | 76.60\% | 29.05\% | 0.018 | 88.34\% | 11.74\% |

## SA1.2 Choice of base year and interaction effects

As a robustness check and to account for interaction effects in the counterfactual analysis, we reverse the role of the base year and the target year in our reweighting procedure. So far, we have considered the wage distribution in 2010 and changed the distribution of characteristics back to that of the base year 1985. This is indicative of the part of the inequality increase that could be 'reversed' by undoing the change in characteristics. In this case, the inequality change explained by composition effects is $Q G\left(t_{w}=2010, t_{x}=2010\right)-Q G\left(t_{w}=2010, t_{x}=1985\right)$. Now, we focus on the opposite case in which we start with the wage distribution in 1985 but only change the distribution of characteristics to the level of 2010. This correspondends to the change $Q G\left(t_{w}=1985, t_{x}=2010\right)-Q G\left(t_{w}=1985, t_{x}=1985\right)$, i.e. the part of the inequality increase that can be accounted for by solely changing the distribution of characteristics while holding fixed the conditional wage structure of 1985.

Figures SA5 to SA8 and tables SA6, SA7 report the findings. For males, the contribution of the different sets of covariates to the overall inequality increase remain qualitatively similar, with a few notable exceptions. The general result is that compositional changes in educational qualifications and in labor market histories provide substantial contributions, while compositional changes related to potential work experience and the occupations/industry structure do so only to a much smaller extent (table SA6 vs. table 4). However, the impact of education changes is much stronger in table SA6 compared to table 4 ( $31.9 \%$. $59.4 \%, 7.4 \%, 18.6 \%$ vs. $17.1 \%, 37.5 \%,-1.0 \%, 7.1 \%$ ). This means that compositional changes over time are associated with a stronger rise in wage inequality based on the wage distribution in 1985 compared to 2010. ${ }^{17}$ Put differently, the effects of a widening conditional wage structure $f(w \mid x)$ is stronger when applied to the distribution of characteristics in 1985 than when applied to that in 2010. This would naturally arise if the 1985 distribution of characteristics is more heterogeneous so that applying diverging wage returns to this more heterogeneous population leads to stronger inequality increases. Take education, the share of low-skilled declines from a high initial level, while the share of high-skilled increases (figure 7). Another difference between tables 4 and

[^0]is equivalent to
$$
Q G\left(t_{w}=10, t_{x}=10\right)-Q G\left(t_{w}=85, t_{x}=10\right)<Q G\left(t_{w}=10, t_{x}=85\right)-Q G\left(t_{w}=85, t_{x}=85\right) .
$$

SA6 is that the contribution of occupations/industries falls when the base year 2010 is used (table SA6). In contrast to the results for education, the composition of occupation and industry has changed in a way that wage inequality increases more strongly for the 2010 composition of occupation and industry compared to the 1985 composition.

For females, the contribution of composition changes in work experience and recent labor market histories remains qualitatively unchanged when we change the base year (columns 6 to 10 in tables 5 and SA7). As for males, the compositional effects of educational upgrading becomes much stronger in table SA7. The only other effect for females, that is not fully robust to the choice of the base year, concerns the changes in occupations and industries. Here, table SA7 shows pronounced effects on inequality in the upper and lower part of the distribution, which are not present in table 5 . The overall contribution of compositional effects to rising female wage inequality in table SA7 is even larger than for the base year 1985 (table 5). In particular, composition changes can account for $78.4 \%$ ( $103.2 \%$ ) of the rise in female overall (lower tail) wage inequality between 1985 and 2010. We conclude that the composition changes would have been associated with a large increase in inequality based on 1985 wages compared to 2010 wages. This is in contrast to the widely held view in the past that Germany used to be a country where institutions strongly limited wage inequality (see Fitzenberger 1999 or Dustmann et al. 2014 for a critical assessment of this view).

Figure SA5: Inequality development base year 2010, specification E


Figure SA6: Inequality development base year 2010, specification EE


Figure SA7: Inequality development base year 2010, specification EEH


Figure SA8: Inequality development base year 2010, specification EEHOI


Table SA6: Reweighted inequality increase 1985-2010, males, compositions for base year 2010

|  | Observed | Ed |  |  | Ed+Ex |  |  | Ed+Ex+Hist |  |  | Ed+Ex+Hist+Occ+Ind |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Increase | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment |
| 85/15 | 0.290 | 0.198 | 31.87\% | 31.87\% | 0.192 | 33.87\% | 2.00\% | 0.153 | 47.34\% | 13.47\% | 0.167 | 42.46\% | -4.88\% |
| 85/50 | 0.137 | 0.055 | 59.40\% | 59.40\% | 0.039 | 71.47\% | 12.07\% | 0.027 | 80.17\% | 8.70\% | 0.041 | 69.80\% | -10.37\% |
| 50/15 | 0.154 | 0.142 | 7.43\% | 7.43\% | 0.153 | 0.48\% | -6.95\% | 0.126 | 18.18\% | 17.70\% | 0.126 | 18.18\% | 0.00\% |
| 90/10 <br> residual | 0.183 | 0.149 | 18.64\% | 18.64\% | 0.138 | 24.62\% | 5.98\% | 0.107 | 41.36\% | 16.74\% | 0.094 | 48.73\% | 7.37\% |

Table SA7: Reweighted inequality increase 1985-2010, females, compositions for base year 2010

|  | Observed |  | Ed |  |  | Ed+Ex |  |  | d+Ex+His |  | Ed+ | x+Hist+O | +Ind |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Increase | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment | Remaining increase | Explained share | Increment |
| 85/15 | 0.218 | 0.159 | 27.10\% | 27.10\% | 0.099 | 54.49\% | 27.39\% | 0.062 | 71.79\% | 17.30\% | 0.047 | 78.39\% | 6.60\% |
| 85/50 | 0.086 | 0.060 | 30.58\% | 30.58\% | 0.044 | 48.38\% | 17.80\% | 0.022 | 74.01\% | 25.63\% | 0.051 | 40.09\% | -33.92\% |
| 50/15 | 0.132 | 0.099 | 24.85\% | 24.85\% | 0.055 | 58.44\% | 33.59\% | 0.039 | 70.36\% | 11.92\% | -0.004 | 103.22\% | 32.86\% |
| 90/10 residual | 0.185 | 0.180 | 2.55\% | 2.55\% | 0.093 | 49.56\% | 47.01\% | 0.061 | 66.94\% | 17.38\% | 0.046 | 75.11\% | 8.17\% |


[^0]:    ${ }^{17}$ This conclusion is based on the following formal argument $(10 \equiv 2010,85 \equiv 1985)$ :

    $$
    Q G\left(t_{w}=85, t_{x}=10\right)-Q G\left(t_{w}=85, t_{x}=85\right)>Q G\left(t_{w}=10, t_{x}=10\right)-Q G\left(t_{w}=10, t_{x}=85\right)
    $$

